## Data 102 Spring 2021

## Midterm 2

- Please write your solutions using either pen/pencil and paper, or a tablet. Each question should start on a new page. At the end of the exam period (or earlier), please upload your exam to the "Midterm 2" assignment on Gradescope. It is your responsibility to make sure your work will be legible!
- We will not answer any questions during the exam. If you think a question is unclear, state your assumptions and answer accordingly.
- You have 80 minutes to work on the exam: you must stop working at 11:00AM PT.
- This exam has 6 questions, for a total of 40 points. You must complete all 6 questions to receive full credit. There are multiple versions of this exam.
- Unless otherwise stated, you must show your work to receive full credit.
- You may, without proof, use theorems and facts that were given in the lectures, homework, lab, or discussions.
- You must complete this honor pledge in order to receive credit on the exam: We ask that you act in accordance with the honor code. Please copy the following statement by hand and sign your name, and include this in your submission.

As a member of the UC Berkeley community, I act with honesty, integrity, and respect for others. These answers are my own work.

0. Make sure you complete the honor pledge on the previous page.

## 1. (5 points) Nonparametric methods.

- (a) (1 point) (True/False) When training a random forest, each tree is trained with the same features, but in a different order.
- (b) (1 point) (True/False) Techniques like LIME use a simple, interpretable model to approximate a more complex model.
- (c) (1 point) (True/False) If there are nonlinear interactions between the input variables and a binary output label, then there is no way to use logistic regression to model the relationship between them.
- (d) (1 point) (True/False) Backpropagation is an algorithm that is only used for training neural networks.
- (e) (1 point) (True/False) When thinking about the bias-variance tradeoff, logistic regression has higher bias than a decision tree (with no depth limit).

2. (15 points) Causal inference. You are working with a developmental economist to understand the effect of free school lunches on school attendance. To study this, the economist conducted a completely randomized experiment that randomized Grade 5 students into receiving a free lunch (T=1) or not (T=0), and then observed whether they attended school (Y=1) or not (Y=0). The results are reported in the following table.

	T = 0	T=1
Y = 0	100	50
Y=1	700	450

Table 1: Grade 5 students

- (a) (2 points) Compute the Neyman (difference-in-means) estimate for the average treatment effect (ATE) of school lunches on school attendance for Grade 5 students.
- (b) (1 point) Write one sentence of plain English interpreting the ATE. Your answer should be understandable to a general audience, and should make the strongest valid conclusion that you can. *Hint: What is the effect of receiving a free school lunch?*
- (c) (1 point) We compute a 95% confidence interval for the true ATE using the Neyman variance. If the interval does not contain 0, which of the following null hypotheses can we reject (at the 95% level)? Select <u>all</u> that apply (or write "none").
  - A. Fisher's strong null hypothesis
  - B. Neyman's weak null hypothesis
- (d) (2 points) The economist simultaneously did a completely randomized experiment on Grade 6 students, with the results reported in Table 2.

	T = 0	T=1
Y = 0	200	200
Y=1	300	300

Table 2: Grade 6 students

For the rest of this question, we investigate the results using a super-population framework. We introduce a covariate X such that X = 1 for students in Grade 6, and X = 0 for students in Grade 5. Compute the estimated propensity score function  $\hat{e}(x)$  for x = 0, 1.

- (e) (2 points) Is X (which grade a student is in) a confounding variable? In one sentence, explain why or why not.
- (f) (2 points) Does the unconfoundedness assumption hold? In one sentence, explain why or why not.

(g) (2 points) The next two parts are about the inverse-propensity weighting (IPW) estimate for the average treatment effect (ATE) of school lunches on school attendance for the **combined population** of Grade 5 and 6 students. The estimate has the form.

$$\hat{\tau}_{IPW} = \frac{1}{n} \left( \frac{A}{\hat{e}(0)} + \frac{B}{\hat{e}(1)} - \frac{C}{1 - \hat{e}(0)} - \frac{D}{1 - \hat{e}(1)} \right). \tag{1}$$

What are the values of A, B, C and D?

- (h) (1 point) What is the value of n in equation (1)?
- (i) (2 points) Denote your answer in (a) using  $\hat{\tau}_5$ , and denote the Neyman estimate for the corresponding ATE computed over Table 2 using  $\hat{\tau}_6$ . The economist proposes four estimates for the average treatment effect (ATE) of school lunches on school attendance for the **combined population** of Grade 5 and 6 students. They are as follows.
  - (A) The IPW estimate from part (g) (Equation (1)).
  - (B) Add up the counts in Tables 1 and 2 and compute the Neyman estimate for the ATE using the resulting table.
  - (C)  $\frac{1}{2}\hat{\tau}_5 + \frac{1}{2}\hat{\tau}_6$ .
  - (D)  $(1 w)\hat{\tau}_5 + w\hat{\tau}_6$ , where  $w = \mathbb{P}(X = 1)$ .

Which of these estimates are unbiased for the true ATE? Select <u>all</u> that apply (or write "none").

## 3. (3 points) Instrumental variables.

Consider the linear structural model

$$Y = \alpha + \tau Z + \beta X + \epsilon,$$

$$Z = \alpha' + \gamma W + \beta' X + \delta.$$

We wish to estimate the treatment effect  $\tau$  of Z on Y using W as an instrumental variable. In order for W to be valid instrumental variable, we need some assumptions on Y, Z and W and X. For each of the quantities below, specify whether it **must be zero** (= 0), **must be nonzero** ( $\neq$  0), or **does not matter**.

- (i) Cov(W, Y)
- (ii) Cov(W, X)
- (iii) Cov(W, Z)
- (iv)  $Cov(W, \epsilon)$
- (v)  $Cov(W, \delta)$

- 4. (4 points) Concentration inequalities. Let  $X_1, X_2, \ldots, X_n$  be independent, identically distributed (i.i.d.) random variables, each with mean 0, and having the same distribution as a  $\sigma$ -sub-Gaussian random variable X. Let  $S_n = \sum_{i=1}^n X_i$ .
  - (a) (2 points) Suppose we are told (only for this part of the question) that X is bounded between -a and a. Based on this information, what is a valid value for  $\sigma^2$ ? State the smallest possible valid value.
  - (b) (2 points) By Hoeffding's inequality, we have

$$\mathbb{P}(|S_n| > t) \le \exp\left(-\frac{t^2}{2n\sigma^2}\right). \tag{2}$$

Which of the following changes can we **make** to the assumptions and still guarantee that inequality (2) still hold? **Select all that apply (or write "none").** 

- (A)  $X_1, \ldots, X_n$  are not identically distributed.
- (B)  $X_1, \ldots, X_n$  are not independent.
- (C) We have  $\mathbb{E}[X_i] = \mu_i$  (not necessarily 0) for each  $i = 1, \ldots, n$ , but  $\sum_{i=1}^n \mu_i = 0$ .
- (D) Each  $X_i$  is  $\sigma_i$ -sub-Gaussian (not necessarily all the same), with  $\sum_{i=1}^n \sigma_i^2 = n\sigma^2$ .

5. (8 points) **Bandit Algorithms.** Consider a bandit environment with K=2 arms, with 1-sub-Gaussian arm reward distributions  $P_a$  and means  $\mu_a$  for a=1,2. Assume that arm 1 is the optimal arm (i.e.,  $\mu_1 > \mu_2$ ), and so we may define the suboptimality gap  $\Delta = \mu_1 - \mu_2$ .

A learner has already played 7 rounds. We are told they pulled arm 1 a total of 2 times, and arm 2 a total of 5 times.

- (a) (2 points) Which of the following are possible policies that the learner was following? Select all that apply (or write "none").
  - (A) The upper confidence bound algorithm (UCB).
  - (B) Explore-then-commit (ETC) with 3 rounds of exploration (m=3).
  - (C) Thompson-sampling (TS).
- (b) (2 points) Suppose the learner was following a deterministic strategy (i.e. the arm choices were determined before the start of the algorithm. We still assume arms 1 and 2 were pulled 2 and 5 times respectively.) As usual, denote the observed average reward for each arm by  $\hat{\mu}_a = \frac{1}{T_a(7)} \sum_{s=1}^7 X_s \mathbf{1}(A_s = a)$  for a = 1, 2. Using Hoeffding's inequality, we compute the following probability bound:

$$\mathbb{P}(\hat{\mu}_1 - \hat{\mu}_2 > t) \le \exp\left(-\frac{(t+A)^2}{B}\right).$$

What are the values for A and B? Express your answer in terms of  $\mu_1, \mu_2$ , and  $\Delta$ .

(c) (2 points) Suppose we were running Thompson Sampling with Gaussian priors and likelihoods. For concreteness, suppose  $\mu_1 = 5$ ,  $\mu_2 = 3$ ,  $\hat{\mu}_1(7) = 4.1$  and  $\hat{\mu}_2(7) = 2.5$ . Let  $\mathcal{N}(z_a, v_a^2)$  be the prior for arm a = 1, 2. For each of the following parameters, we have suggested a number of possible values. For each parameter, choose the **one** option that maximizes the probability of pulling arm 1 in the next (eighth) round.

$$z_1:$$
 -10, 4.1, 5, 10  
 $z_2:$  -10, 2.5, 3, 10  
 $v_1^2:$  0.05, 10  
 $v_2^2:$  0.05, 10

For example, if you chose a value of 10 for all parameters, your solution might look like:  $z_1 = 10$ ,  $z_2 = 10$ ,  $v_1^2 = 10$ ,  $v_2^2 = 10$ . (This is not necessarily the correct answer, just an example.)

(d) (2 points) A very risk-averse data scientist has proposed the following *lower* confidence bound algorithm. Define the lower confidence bound

$$LCB_a(t, \delta) = \hat{\mu}_a(t) - \sqrt{\frac{2\log(1/\delta)}{T_a(t)}}$$

At each round t, the learner selects:

$$A_t = \begin{cases} t & t \le K \\ \operatorname{argmax}_{a=1,\dots,K} \operatorname{LCB}_a(t-1, 1/t^3) & t > K. \end{cases}$$

Does this algorithm have logarithmic regret? Explain why or why not. You don't need to provide a full proof, but you must provide a convincing explanation.

6. (5 points) Uncertainty quantification for GLM. You are consulting for an ice-cream company that wants to investigate the relationship between mean daily temperature X (in degrees Celsius) and the number of ice-cream cones sold Y (a count). You model this using Poisson regression, with

$$\mathbb{E}[Y|X] = e^{\beta_0 + \beta_1 X}. (3)$$

To fit the model, we use a data set S that contains i.i.d. samples  $(X_1, Y_1), \ldots, (X_n, Y_n)$  from our population of interest, obtaining coefficient estimates  $\hat{\beta}_0$  and  $\hat{\beta}_1$ .

- (a) (2 points) We wish to use the bootstrap to get a 95% confidence interval for  $\beta_1$ . We have already generated 1000 bootstrap replicates of  $\hat{\beta}_1$ , which are stored in a one-dimensional numpy array beta\_boot. Write no more than two lines of code in Python that gives the left and right end-points of such an interval. You may assume that we have already run the line import numpy as np.
- (b) (1 point) After running your code, you discover a bug: each bootstrap replicate  $\hat{\beta}^*$  was obtained by drawing 2n samples at random with replacement from S (instead of just n samples). Compared to the correct bootstrap confidence interval, is the width of your confidence interval smaller, larger, or roughly the same?
- (c) (1 point) Suppose we know that the mean temperature tomorrow is going to be 35°C. Given a model with regression coefficients  $\beta = (\beta_0, \beta_1)$ , what is the probability  $p(Y = 90|X = 35, \beta)$  that 90 ice-cream cones will be sold tomorrow?
- (d) (1 point) Suppose we instead use a Bayesian approach, and fit a Bayesian Poisson GLM. Let  $q(\beta)$  denote the posterior distribution (density) that we compute over  $\beta = (\beta_0, \beta_1)$ . Write a formula for the **posterior predictive probability** that 90 ice-cream cones will be sold tomorrow given that the mean temperature tomorrow is going to be 35°C. You may leave your answer in terms of q and  $p(Y = 90|X = 35, \beta)$ .