

# Lecture 12: Robust Uncertainty via the Bootstrap

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# Announcements

- Congratulations on finishing Midterm 1!!
- Vitamin out today, due Sunday
- HW3 out today, due in 2 weeks
- Lab and discussion resume as normal next week

# Recap

- Bayesian, frequentist regression
- Model mis-specification  $\implies$  overly narrow uncertainty (for both Bayesian and frequentist)
- Model checking

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This time: more robust frequentist uncertainty estimates, via bootstrap

# Bayesian vs. frequentist uncertainty

**Credible interval:** Posterior probability that  $\theta$  lies in interval is at least  $p$

**Confidence interval:** Conditional on  $\theta$ , interval contains  $\theta$  with probability  $p$

# Bayesian vs. frequentist uncertainty

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**Confidence interval:** Conditional on  $\theta$ , interval contains  $\theta$  with probability  $p$

- Another interpretation: no matter what the true parameters are, interval contains them 99% of the time (for  $p = 0.99$ )
- This property is called *coverage*

# Confidence vs. credible intervals

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- If true  $\theta$  has low prior probability, might not have coverage

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Why is a confidence interval not (necessarily) a valid credible interval?

- Suppose you observe 6 coin flips that all come up heads, but you have very high prior probability that the coin is fair. The 95% confidence interval won't contain  $\frac{1}{2}$ , but the credible interval should.



# Confidence vs. credible intervals

Other distinction: source of randomness

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Confidence interval requires imagining hypothetical “other” draws of data. We’ll see this used later for the bootstrap.

We’ll focus on confidence intervals for the rest of this lecture.

# Confidence intervals for regression

Recall wind turbines example:

$$\mathbb{E}[\text{Turbines} \mid \text{Year}] = \exp(\beta_{\text{Year}} \cdot \text{Year} + \beta_{\text{Intercept}})$$

To understand growth rate, care about coefficient  $\beta_{\text{Year}}$

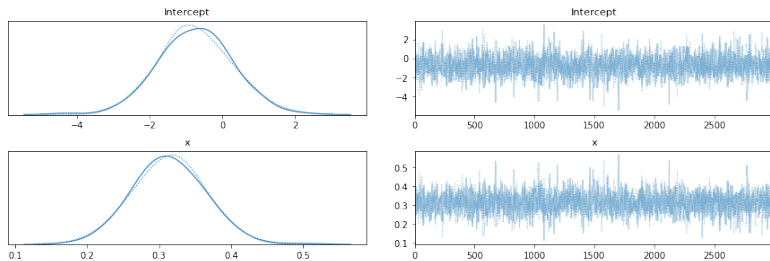
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Previously: MCMC sampling gives us posterior distribution (and hence credible interval) for  $\beta_{\text{Year}}$ :



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General recipe: use generalization of CLT called “asymptotic normality”

Beyond scope of this class, but `statsmodels` package will do it for us!

# Confidence intervals with statsmodels

[Jupyter demo]



# Escaping model mis-specification

Frequentist confidence intervals can be wrong if model is wrong

- Just like Bayesian case

We'll escape this with a **non-parametric** tool for producing frequentist CIs

Non-parametric  $\implies$  doesn't rely on model  $\implies$  more robust

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You've seen this before: the **bootstrap**

# The Bootstrap

Idea for computing confidence intervals + uncertainty

Without bootstrap:

- Chi-square test, student-t test, . . .
- Lots of algebra, need different formula for each setting
- Often rely on model assumptions

With bootstrap:

- Single unified approach
- Computer simulation
- Fewer assumptions

# Demo: Mean Estimation

[Jupyter demo]

# Bootstrap: formal setting

Data:  $x_1, \dots, x_n$

Estimator:  $\hat{\theta} = \hat{\theta}(x_1, \dots, x_n)$

- $\theta^*$ : population parameter (what  $\hat{\theta}$  converges to as  $n \rightarrow \infty$ )

Question: How close is  $\theta^*$  to  $\hat{\theta}$ ?

## Some concrete examples

Mean of 1-dimensional distribution:

- $x_1, \dots, x_n \in \mathbb{R}$
- $\hat{\theta}(x_1, \dots, x_n) = \frac{1}{n}(x_1 + \dots + x_n)$

How close is estimate to the true mean?

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Regression:

- $(x_1, y_1), \dots, (x_n, y_n) \in \mathbb{R}^d \times \mathbb{R}$
- $\hat{\beta}((x_1, y_1), \dots, (x_n, y_n)) = \operatorname{argmin}_{\beta} \sum_{i=1}^n (y_i - \beta^\top x_i)^2$

How close is  $\hat{\beta}$  to population parameters  $\beta^*$ ?

# More complex examples

Mixture models

Density estimation

Neural nets? (Actually not...)



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Population distribution  $p^*$

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Imagine hypothetically sampling fresh data:

$$x_1, \dots, x_n \rightarrow \hat{\theta} \text{ (Original sample)}$$

$$x'_1, \dots, x'_n \rightarrow \hat{\theta}' \text{ (Re-sample)}$$

$$x''_1, \dots, x''_n \rightarrow \hat{\theta}''$$

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⋮

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⋮

Implicit commitment: distribution of  $\hat{\theta}$  roughly centered on  $\theta^*$  (low bias)

# The Bootstrap

Want to approximate hypothetical samples  $\hat{\theta}', \hat{\theta}'', \dots$

But only have actual data  $x_1, \dots, x_n \rightarrow \hat{\theta}$

Idea: subsample data

- With replacement
- $n$  points in each sample

# Bootstrap: Pseudocode

$B$ : number of bootstrap samples

For  $b = 1, \dots, B$ :

- Sample  $x'_1, \dots, x'_n$  with replacement from  $x_1, \dots, x_n$
- Let  $\hat{\theta}^{(b)} = \hat{\theta}(x'_1, \dots, x'_n)$

Output  $\{\hat{\theta}^{(1)}, \dots, \hat{\theta}^{(B)}\}$

# Bootstrap in python

[Jupyter demos]

# Counterexample: Max

[Jupyter demo]



## Counterexample: Max

$$\hat{\theta}(x_1, \dots, x_n) = \max_{i=1}^n x_i$$

$n$  samples: always finite

$\infty$  samples: infinite

# When does the bootstrap work?

Most parametric estimators are fine

- I.e. fixed number of parameters  $d$  and  $d \ll n$

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- Decision trees
- Neural nets
- Kernel regression

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Other commitments:

- $\hat{\theta}$  approximately unbiased
- $\theta^*$  is a meaningful quantity

# Summary

- Credible intervals vs. confidence intervals
- Confidence intervals in statsmodels
- Still depend on assumptions!
- Bootstrap more robust (and flexible)